The Identification Power of the Markov Assumption in Dynamic Discrete Choice Models

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Identification Under the Markov Property

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Motivation

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Identification Under the Markov Property

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- Suppose we have an important economic question
- To answer this question, we use data on the dynamic behavior of economic agents
- Dynamic discrete choice models are tools to analyze these data
- The Markov assumption is a key part of these models
- Many important implementations
- Many important papers about the properties of these tools

- What testable restrictions does the Markov assumption place on the data?
- What happens if the agent knows more than the researcher?
- **③** What can the data tell us about the unobserved information?

1 Dynamic Discrete Choice Framework with Unobserved Heterogeneity

- The Observed Markov Property
 - Observable Implications of the Markov Assumption
 - Unobserved Heterogeneity that Preserves the Markov Property
 - Unobserved Heterogeneity that Violates the Markov Property
 - Identification of the Unobserved State
- Testing the Markov Property
- Special Cases
 - Persistent Types
 - Stationary Initial State
 - Time Dependent Policy Function
- Monte Carlo Evidence and Applications

"Short" Panel:

- Agents/markets: $i = 1, \ldots, n$
- Time periods: $t = 1, \ldots, T$
- Actions: a_{it}
- Environment: x_{it}
- Asymptotics: $n \rightarrow \infty$, T fixed

Can directly estimate: $P\{(a_t, x_t), t = 1, ..., T\}$

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- Discrete time index: t = 1, 2, ...
- State space: $s_t \in \mathcal{S}$, cardinality is known: $|\mathcal{S}| = S$
- Probability distribution over the initial state: $P(s_1)$
- Action set: $a_t \in \mathcal{A}$, cardinality is known: $|\mathcal{A}| = A$
- State transition probability function: $P_t(s_{t+1}|s_1, a_1, \dots, s_t, a_t)$
- Utility functional: $V(s_t) = \mathbb{E}_t \left[\sum_{\tau=t}^{+\infty} \beta^{\tau-t} v(s_{\tau}, a_{\tau}) + \epsilon_{\tau}(a_{\tau}) \right]$, that consists of:
 - the single period utility function $v(s_t, a_t): \mathcal{S} imes \mathcal{A} o \mathbb{R}$
 - the agent's discount factor $\beta \in ({\tt 0}, {\tt 1})$
 - the idiosyncratic action-specific error term ε_τ = (ε_τ(1),..., ε_τ(A)) drawn independently across time from a distribution defined by a p.d.f: f(ε_t|s_t) : ℝ^A × S → ℝ

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Assumptions

- The state transition $P_t(s_{t+1}|s_t, a_t)$ has the Markov property: $P_t(s_{t+1}|s_1, a_1, \dots, s_t, a_t) = P_t(s_{t+1}|s_t, a_t)$ for any t > 1.
- The state transition is time homogenous: $P_t(s_{t+1}|s_t, a_t) = P(s_{t+1}|s_t, a_t)$ for any t.

Proposition There exists an optimal decision rule that has the Markov property: $d^*(s_t, \epsilon_t) = a_t$

Conditional choice probabilities: $P(a_t|s_t) = \mathbb{E}[\mathcal{I}(d^*(s_t, \epsilon_t) = a_t)].$

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Unobserved Heterogeneity = the researcher has only partial information about the state of the world

Formally,

- $s_t = (x_t, u_t)$
- the agent observes s_t
- the researcher observes x_t but not u_t
- define a coarsening function $U: \mathcal{S} \rightarrow \mathcal{X}$
- denote X = |U(S)|
- unobserved heterogeneity is present when X < S

Data

- The date are defined by $P\left\{(a_t, x_t), t = 1, \dots, T\right\}$
- Its dimensionality: $(AX)^T 1$

Can these data be rationalized by a dynamic discrete choice model that has the Markov property?

- without unobserved heterogeneity
- with unobserved heterogeneity

Assume S = X

Model

- Elements of the Structure:
 - initial state: $P(s_1)$
 - state transition: $P(s_{t+1}|s_t, a_t)$
 - optimal policy: $P(a_t|s_t)$
- Its dimensionality: $(S-1) + (S-1)AS + (A-1)S = AS^2 1$

Thus, the Markov assumption reduces the dimensionality of the object from $(AS)^{T} - 1$ to $AS^{2} - 1$. The number of imposed restrictions grows exponentially with T.

Example

- *S* = 3, *A* = 2, *T* = 5
- dimensionality of the generic process: $6^5 1 = 46,656$
- dimensionality under the Markov assumption: $2 \cdot 3^1 1 = 7$
- Probability of a randomly chosen process to violate the Markov property: 99.985%

Suppose X < S. What can we infer from the data about the unobserved heterogeneity?

Four reasons for lack of identification:

- relabeling
- irrelevance (zero-probability states)
- ovefitting
- collinearity

Unobserved Heterogeneity Overfitting

- Dimensionality of the data: $(AX)^T 1$
- Dimensionality of the model: $AS^2 1$

Proposition 2: If $T < \frac{\log(A) + 2\log(S)}{\log(A) + \log(X)}$, then the primitives of the model are not identified.

Proposition 3: If $X < \frac{S^{2/T}}{A^{\frac{T-1}{T}}}$, then the primitives of the model are not identified.

Proposition 4: Generically, the dimensionality of the data does not equal to the dimensionality of the model.

Thus,

- The panel cannot be too short
- The unobserved heterogeneity cannot be too rich
- The unobserved heterogeneity is either under- or overidentified

Definition Two states s and s' are called collinear if:

- U(s) = U(s')
- P(a|s) = P(a|s') for all $a \in \mathcal{A}$
- $\sum_{\tilde{s} \text{ s.t. } U(\tilde{s})=const} P(\tilde{s}|a,s) = \sum_{\tilde{s} \text{ s.t. } U(\tilde{s})=const} P(\tilde{s}|a,s')$

Proposition 5 The primitives of the model are identified for some T if and only if the state space S does not have collinear states.

Proposition 6 Suppose the data are Markov. Then either there is no unobserved heterogeneity or the state space has collinear states.

Should we be worried about unobserved heterogeneity?

- If the data are Markov, then
 - either there is no unobserved heterogeneity
 - or we cannot tell anything about the unobserved state
- If the data are not Markov, then
 - we can identify unobserved heterogeneity as long as it is not too rich
 - if we can identify unobserved heterogeneity, we will have a set of over identifying restrictions that we can test

Standard testing framework:

- $H_0 =$ the Markov model
- $H_1 =$ unrestricted process

Three tests are available:

- Estimate the Markov model and run a LM test
- Estimate the unrestricted model and run a Wald test
- Estimate both models and run a LR test

- J persistent types: the unobserved part u_t doesn't change over time
- dimensionality of the model:
 - initial state: (X-1)J
 - state transition: (X 1)XJ
 - policy function: (A-1)XJ
 - distribution of types: (J-1)
 - overall: $X^2J + AXJ 1$
- dimensionality of the Data: $(AX)^T 1$
- Threshold $T = \frac{\log(X) + \log(J) + \log(A + X)}{\log(A) + \log(X)}$
- Threshold $J = \frac{A^T X^{T-1}}{A+X}$
- No collinear states if different types either act or affect state transition differently

- $P(S_1)$ is the stationary distribution of the Markov process
- Dimensionality of the model reduces to $(S-1)AS + (A-1)S = AS^2 S$
- Dimensionality of the data: $(AX)^T 1$

• Threshold
$$T = \frac{\log(AS^2 - S + 1)}{\log(A) + \log(X)}$$

• Threshold $X = \frac{(AS^2 - S + 1)^{1/T}}{A}$

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- Suppose the optimal policy is time-dependent (e.g. finite-horizon)
- Dimensionality of the model:
 - initial state: (S-1)
 - state transition: (S-1)SA
 - policy function: (A-1)ST
 - overall: $AS^2 + (T-1)S(A-1) 1$
- Dimensionality of the data: $(AX)^T 1$
- Dimensionality of the model grows linearly in T
- Dimensionality of the data grows exponentially in T
- Therefore, the data are still informative

- Unlike many econometric models, the Markov assumption imposes a number of *testable* restrictions on the data.
- If the data do not reject the Markov property, we cannot say much about unobserved heterogeneity.
- If the data reject the Markov property, then unobserved heterogeneity can be potentially recovered from the data.

- Monte Carlo properties of different tests
- What can we say about unobserved heterogeneity in the well-known applications (e.g. Rust (1987), Ryan (2012))?
- What exactly will we estimate if a part of the state is unobserved but the Markov property still holds?
- What exactly will we estimate if a part of the state is unobserved and the Markov property fails?